Breeding goal traits accounting for feed intake capacity and roughage or concentrate intake separately

Margreet Heida, Ghyslaine C. B. Schopen, Marinus F. W. te Pas, Birgit Gredler-Grandl,
and Roel F. Veerkamp*

Wageningen University & Research Animal Breeding and Genomics, PO Box 338, 6700 AH Wageningen, the Netherlands

ABSTRACT

Current breeding tools aiming to improve feed efficiency use definitions based on total dry matter intake (DMI); for example, residual feed intake or feed saved. This research aimed to define alternative traits using existing data that differentiate between feed intake capacity and roughage or concentrate intake, and to investigate the phenotypic and genetic relationships among these traits. The data set contained 39,017 weekly milk yield, live weight, and DMI records of 3,164 cows. The 4 defined traits were as follows: (1) Feed intake capacity (FIC), defined as the difference between how much a cow ate and how much she was expected to eat based on diet satiety value and status of the cow (parity and lactation stage); (2) feed saved (FS), defined as the difference between the measured and the predicted DMI, based on the regression of DMI on milk components within experiment; (3) residual roughage intake (RRI), defined as the difference between the measured and the predicted roughage intake, based on the regression of roughage intake on milk components and concentrate intake within experiment; and (4) residual concentrate intake (RCI), defined as the difference between the measured and the predicted concentrate intake, based on the regression of concentrate intake on milk components and roughage intake within experiment. The phenotypic correlations were -0.72 between FIC and FS, -0.84 between FS and RRI, and -0.53 between FS and RCI. Heritability of FIC, FS, RRI, and RCI were estimated to be 0.21, 0.12, 0.15, and 0.03, respectively. The genetic correlations were -0.81 between FS and FIC, -0.96 between FS and RRI, and -0.25 between FS and RCI. Concentrate intake and RCI had low heritability. Genetic correlation between DMI and FIC was 0.98. Although the defined traits had moderate phenotypic correlations, the genetic correlations between DMI, FS, FIC, and RRI were above 0.79 (in

absolute terms), suggesting that these traits are genetically similar. Therefore, selecting for FIC is expected to simply increase DMI and RRI, and there seems to be little advantage in separating concentrate and roughage intake in the genetic evaluation, because measured concentrate intake was determined by the feeding system in our data and not by the genetics of the cow.

Key words: feed efficiency, DMI, feed saved, breeding goal

INTRODUCTION

Feed costs are approximately 50% of total costs on a dairy farm (FADN, 2011). Selection on feed efficiency in dairy cattle can lead to reduced feed costs, as animals will require less feed for milk production. Additionally, selection on feed efficiency could lead to genetic improvement and reduction of methane excretion (Arthur and Herd, 2005; de Haas et al., 2017). Therefore, breeding for feed-efficient cows has become even more important over the past few years.

To breed for feed efficiency, feed intake and production records need to be recorded, and feed efficiency traits need to be defined to estimate (genomic) breeding values for these traits. Subsequently, breeding values can be used to breed for feed efficiency in the required direction. The definition of feed efficiency in dairy cows has undergone scientific debate for a long time (Veerkamp, 1998; Berry and Crowley, 2013; Veerkamp et al., 2013; Pryce et al., 2015; Hurley et al., 2017; Tempelman and Lu, 2020). Major elements of this debate are summarized here. Feed efficiency can be calculated as gross feed efficiency or feed conversion ratio, which is easier to understand for the end user (e.g., farmers). However, these ratio traits favor animals with high milk production rather than net efficient cows, because maintenance costs are diluted (Veerkamp and Emmans, 1995). Furthermore, it is difficult to predict the response to selection for the underlying traits (e.g., feed intake, production traits) and feed efficiency when selection is for the ratio trait (Gunsett, 1984) because of dependencies between the variance and covariance

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*Corresponding author: Roel.Veerkamp@wur.nl

components and the mean of the underlying traits (Veerkamp and Emmans, 1995). Also, improving feed efficiency ratios does not necessarily lead to improvement of energy conversion toward milk (Veerkamp and Emmans, 1995). Alternatively, efficiency has been defined as the linear difference between the observed and the expected feed intake, as is the case for residual feed intake (**RFI**), which is calculated as the residuals of the regression of milk production and other energy sinks on feed intake (Koch et al., 1963; Veerkamp et al., 1995; Li et al., 2020). The RFI is phenotypically independent of its regressors by construction and not affected by the dilution of maintenance costs by increased milk yield, unlike gross efficiency or feed conversion ratio. So far, RFI is the closest approximation of net feed efficiency (Veerkamp et al., 2013), although it might be sensitive to inaccurately modeled factors that could also lead to variation in residuals (Li et al., 2017), or to an imperfect regression model used to calculate RFI. Using RFI has another major disadvantage for dairy cows: RFI is, by construction, independent of (metabolic) weight; therefore, a very heavy cow gets a favorable RFI if she ate less than expected based on her milk production and body weight. In fact, if 2 cows of different body weight eat the same amount and have the same milk production, the heavier cow gets the best RFI as the most net efficient. However, the heavier cow will necessarily allocate more energy to maintenance. A suggested alternative could be feed saved (**FS**), defined as the feed saved when the animal consumes less feed than other animals to produce the same amount of milk (Pryce et al., 2015). For FS, animals are penalized for extra body weight as well. A larger animal gets a less-favorable value for FS compared with a smaller animal. This debate about the definition of efficiency remains lively because of the difficulty of defining the exact sources of genetic variation, which is often investigated with overly limited data (Veerkamp and Emmans, 1995).

Another issue related to the definition of feed efficiency in the definition of current breeding goals is the lack of differentiation between roughage and concentrate intake. Especially now, the aim is to increase the amount of roughage (human-inedible food) in the ration of Dutch dairy cows. Sixty-five percent of dietary protein has to be produced in the Netherlands from the year 2025 onward (Kleijne et al., 2018). This can be achieved by increasing the amount of grass in the ration while reducing maize and concentrate (Hilhorst, 2018). Also, based on nutritional models, it could be argued that cows that eat more roughage per kilogram of milk can be fed less concentrate and can also convert more grass into milk (Groen and Korver, 1989; Harrison et al., 1990; Veerkamp, 1998; Buckley et al., 2005). This hypothesis that cows should eat more roughage is also

used by breed-oriented breeders who prefer big and heavy cows. The answer to the question of whether we should breed for cows that can eat more roughage is partly economic and depends on the cost of roughage versus concentrate (Koenen et al., 2000).

Nutritional model predictions ignore genetic variation between animals and the ability to reduce yield production on a given diet, which has been shown in classic genotype by environment interaction (**G**×**E**) experiments (Pryce et al., 1999; Coffey et al., 2004). Also, Veerkamp et al. (2018) investigated whether we should select for cows that can eat more or cows that are more efficient, from the viewpoint of **G**×**E**. In that study, DMI data were divided between high concentrate versus high roughage-based rations, and heritability and genetic correlations were estimated for these 2 distinct traits, which can be seen as 2 different environments. Then selection responses were predicted for either a high concentrate or high roughage-based ration. There was too little evidence for **G**×**E** to suggest that selecting for increased DMI on the high roughage ration is beneficial for performance on the high roughage-based ration. The study presented here addresses the same question about concentrate versus roughage efficiency but from a different viewpoint. Rather than investigating **G**×**E** for total DMI, the objective here was to use existing data, which is used in the national genetic evaluations for DMI, to define alternative traits related to feed intake capacity and roughage or concentrate intake separately and quantify their phenotypic and genetic parameters, an approach that was also used by Tarekegn et al. (2021). Insight on these novel traits might help to judge whether it is worthwhile to separate these traits in the genetic evaluation, to breed for intake capacity or to breed for concentrate- or roughage-efficient cows separately.

MATERIALS AND METHODS

Data Description

The data used in this study were a subset of the data used for the national DMI genetic evaluations in the Netherlands and were collected during several short- and long-term experiments on farms of Wageningen University & Research in the Netherlands [Aver Heino (Heino), Bosma Zathe (Ureterp), Cranendonck (Soerendonck), t Gen (Lelystad), Minderhoudhoeve (Swifterbant), Waiboerhoeve Dairy Unit 2 and 3 (Lelystad), Zegveld Farm (Zegveld), Hoorn (Lelystad), New Wairboerhoeve (Lelystad), and Dairy Campus (Leeuwarden)]. A description of the methodology of most experiments has been summarized previously (Veerkamp et al., 2000; Beerda et al., 2007; Zom et al., 2012; Man-

zanilla Pech et al., 2014; van Kneegsel et al., 2014). All cows were kept indoors in conventional cubicle housings and were offered partial mixed diets ad libitum, and concentrate was given in separate dispensers on an individual-cow basis. The amount of concentrate dispensed varied by each experiment, reflecting common practices in the Netherlands. Cows were milked twice per day, except for 50 cows at Bosma Zathe ($n = 50$), which were milked 3 times per day. Individual animal DMI were transformed to weekly measurements by averaging over 7-d (or sometimes 5-d) measurements. Roughage DMI (**DMIR**) was the partial mixed ration fed to the cows (mainly grass and maize silage, but often complemented with other products), whereas concentrate DMI (**DMIC**) was the amount of concentrate fed separately in the concentrate dispenser or milking parlor. Only those weekly DMI records were kept that had milk yield, fat and protein production, and cow live weight (**LW**) measured in the same week.

Experiments with a concentrate portion of 1 kg/d or less were discarded because a TMR was fed to these cows. After excluding extreme experiments and deleting records with milk yield below 5 kg of milk/day, 0.3 kg of fat/day, and 0.3 kg of protein/day, the new traits were calculated. For the analysis of the traits, records between ± 3 standard deviations (SD) from the mean were retained for all traits. The final data set (Table 1) contained 39,017 records from 3,164 cows of 789 sires and 2,471 dams, and 1,073 dams also had records themselves. Sires had on average 3.9 daughters, ranging from 1 to 54 daughters per sire.

Definition of Traits

Four different traits were defined: feed intake capacity (**FIC**), FS, residual roughage intake (**RRI**), and residual concentrate intake (**RCI**). With FIC, we aimed to measure how much a cow ate in relation to how much she was expected to eat based on her diet: $FIC = DMI_{\text{measured}} - DMI_{\text{expected}}$. A high value for FIC means that the cow ate more dry matter than expected

based on the dairy cow nutrition model (Zom, 2014). The DMI_{expected} was calculated as the expected feed intake (EFI), divided by the satiety value (SV) of that diet: $DMI_{\text{expected}} = \text{EFI}/\text{SV}$. The expected feed intake is calculated as follows (Zom, 2014):

$$\text{EFI} = \left[8.743 + 3.563 \times \left(1 - e^{-1.140 \times a} \right) \right] \times e^{0.1365 \times \left(1 - e^{-0.05889 \times d} \right)} \times \left(1 - 0.05529 \times \frac{g}{220} \right),$$

where a combines the parity and the DIM, $a = \text{parity} - 1 + d/365$, d represents DIM, e represents the exponential constant, and g represents the days that the animal is pregnant. Insemination data were not available in this data set; therefore, it was assumed that cows were pregnant from d 130 in lactation onward. This assumption was based on the average calving interval in the Netherlands, 410 d (CRV, 2019), and the duration of pregnancy, 280 d.

The satiety value (Zom, 2014) of the ration was calculated here as

$$\text{SV} = c_1 \times \text{percentage roughage} + c_2 \times \text{percentage concentrates.}$$

Percentage roughage or concentrate means the percentage of roughage or concentrate DM per DM of the diet; c_1 and c_2 represent diet-specific constants based on, for example, energy and fiber content. Specific information on the c_1 and c_2 for each diet in each experiment was not available, except for amounts of roughage and concentrate intake. Therefore, constants were taken from the national nutritional guidelines for maize and grass silage ($c_1 = 1$) and for concentrates ($c_2 = 0.34$; CVB, 2019).

Feed saved was defined as the difference between the measured DMI and the expected DMI based on fat and protein yield. Unlike definitions of RFI, it was not regressed on LW, and a positive value is favorable. Feed saved was defined as the negative residuals of the following regression:

$$DMI_{ij} = \text{exptreat}_j \cdot \beta_0 + \text{exptreat}_j \cdot \beta_1 \text{kgfat}_i + \text{exptreat}_j \cdot \beta_2 \text{kgprotein}_i + e_{ij},$$

where DMI_{ij} was the dry matter intake of record i , and β_0 , β_1 , and β_2 represented the intercept and the regression coefficients for kilograms of fat and kilograms of protein of record i , respectively (kgfat_i and kgprotein_i), within experimental treatment j (exptreat_j) to allow for different intercepts and slopes among experimental

Table 1. General information on the edited data set

Item	No.
Experimental treatments in data	517
Total number of weekly records for each trait	39,017
Unique animals in data	3,164
Total number of animal lactations	5,209
Average number of records per experimental treatment	75.5
Average number of animals per experimental treatment	14.18
Average number of weekly records per animal	12.3
Average number of daughters per sire	3.9

treatments j , and e_{ij} stands for the random residual effect. For FS we calculated the residual as the difference between fitted value and measured value. A high value for FS means that animals ate less than expected based on their fat and protein production.

Residual roughage intake was defined as the differences between the measured roughage DMI and the expected roughage DMI. Like FS, the expectation was based on the milk production level, but additionally the amount of concentrate fed was adjusted for in the model.

Residual roughage intake was defined as the residuals of the following model:

$$\text{DMIR}_{ij} = \text{exptreat}_j \cdot \beta_0 + \text{exptreat}_j \cdot \beta_1 \text{kgprotein}_i \\ + \text{exptreat}_j \cdot \beta_2 \text{kgfat}_i + \beta_3 \text{DMIC}_i + e_{ij},$$

where DMIR_{ij} was the DM roughage intake in kilograms for record i , and β_0 , β_1 , and β_2 represented the intercept and the regression coefficients for kilograms of fat and kilograms of protein of record i , respectively (kgfat_i and kgprotein_i), within experimental treatment j (exptreat_j), to allow for different intercepts and slopes among experimental treatments j , and e_{ij} stands for the random residual effect. Regression coefficient β_3 represents the regression coefficient on concentrate intake (DMIC) of record i . For RRI, we calculated the difference between the measured value and fitted value. A low value for RRI means that the animal ate less roughage than expected based on its concentrate intake and fat and protein production on an average diet.

Residual concentrate intake was defined as the difference between the observed concentrate intake and the expected concentrate intake. Residual concentrate intake was defined similarly to FS and RRI. The DMIC_{ij} was based on the milk components and the amount of roughage fed.

Residual concentrate intake is represented as the residuals of the following model:

$$\text{DMIC}_{ij} = \text{exptreat}_j \cdot \beta_0 + \text{exptreat}_j \cdot \beta_1 \text{kgprotein}_i \\ + \text{exptreat}_j \cdot \beta_2 \text{kgfat}_i + \beta_3 \text{DMIR}_i + e_{ij},$$

where DMIC_{ij} was the DMI of concentrate for record i in kilograms; β_0 , β_1 , and β_2 represented the intercept and the regression coefficients for kilograms of fat and kilograms of protein of record i , respectively (kgprotein_i and kgfat_i), within experimental treatment j (exptreat_j) to allow for different intercepts and slopes among experimental treatments j . Regression coefficient β_3 represents the regression coefficient on roughage intake (DMIR) of record i , and e_{ij} represents the random

residual effect. For RRI we calculated the difference between the measured value and fitted value. A low value for RCI means that animals ate less concentrate than expected based on their roughage intake and milk production.

Statistical Analysis

Initially, an exploratory analysis was performed for the 4 defined traits FIC, FS, RRI, and RCI using simple descriptive statistics (histograms, Pearson correlations, and scatterplots) using R (R Core Team, 2019). Associations with milk yield, LW, and DMI were visualized by calculating the average (and standard error, SE) for the 4 defined traits, based on bins of 1 kg of milk or DMI and 10 kg of LW, respectively. Also, the mean and SE for the 4 defined traits were calculated for parity and lactation stage.

The phenotypic associations between the 4 defined traits and milk yield, LW, and DMI were also estimated using the linear model described for the genetic analysis with a linear regression coefficient on either LW, milk yield, or DMI.

For the genetic analysis, we were interested in variance components for 10 traits: milk yield (**MY**), fat plus protein yield, LW, DMI, DMIC, DMIR, and the 4 defined traits FIC, FS, RRI, and RCI. Phenotypic, genetic, and permanent environmental variances were estimated univariately with REML using a linear mixed model (Gilmour et al., 2015). The following model was used:

$$y_{ijklmnopqr} = \mu + P_j + \beta_1 \times \text{agec}_k(P_j) + \text{WIM}_l \\ + \text{exptreat}_m + \beta_2 \times \text{breedHO}_n + \beta_3 \times \text{breedFH}_o + \beta_4 \\ \times \text{breedMO}_p + \beta_5 \times \text{breedFV}_q + \beta_6 \times \text{breedMRY}_r \\ + a_i + \text{perm}_i + e_{ijklmnopqr},$$

where $y_{ijklmnopqr}$ represents the phenotype for the trait analyzed for individual i , μ was the overall mean, and P_j represented the fixed effect of parity j . Cows in parity 4 or above (up to 11) were grouped in a common class. $\beta_1 \times \text{agec}_k(P_j)$ was the fixed quadratic regression of age at calving in month k nested within parity j , with β_1 as regression coefficient. WIM_l was the fixed effect of lactation stage with week in milk l (1–53 wk in lactation); exptreat_m was the fixed effect of experimental treatment m (518 levels). When the record was not in a specific experiment, a fixed effect of herd-year-season was fitted instead. Four seasons were defined: January to March, April to June, July to September, and October to December. $\beta_2 \times \text{breedHO}_n$, $\beta_3 \times \text{breedFH}_o$, $\beta_4 \times$

breedMO_p, $\beta_5 \times \text{breedFV}_q$, and $\beta_6 \times \text{breedMRY}_r$ were the fixed quadratic regressions of breed proportions n, o, p, q, and r of Holstein, Friesian Holstein, Montbéliarde, Fleckvieh, and Meuse-Rhine-Issel, respectively, with β_2 to β_6 being the corresponding regression coefficients. a_i was the random additive genetic effect of animal and $\text{var}(a) = \mathbf{A}\sigma_a^2$, where $\mathbf{A}$ was the additive genetic relationship matrix and σ_a^2 the additive genetic variance. Perm_i was the permanent environmental effect nested within parity o, and $\text{var}(\text{perm}) = \mathbf{I}\sigma_{\text{perm}}^2$ was the permanent environmental variance, for perm_i ; there was no grouping of parities higher than 4. Finally, $e_{ijklmnopqr}$ was the random residual term, with $\text{var}(e) = \mathbf{I}\sigma_e^2$.

Estimating genetic and phenotypic correlations between the 10 traits simultaneously is difficult, especially when the traits are linear combinations of each other (for example, DMI is the sum of DMIR and DMIC) and when the SE of the estimated correlations are relatively large. Dependencies between the traits and large SE make the (co-)variance matrices more prone to become nonpositive definite, which might give convergence problems for AI-REML algorithms. Therefore, several approaches were used to estimate the correlation, using the model described and including genetic, permanent environmental, and residual covariances between the traits. Initially 45 bivariate analysis were performed in ASReml (Gilmour et al., 2015) that all converged eventually. Second, the full 10-trait model was started in ASREML version 4.2nc (Jan. 20, 2021), using a combination of AI and (PX)EM-REML algorithms. After more than 600 iterations in total, parameters remained stable, and the log-likelihood changed with 0.08 only in the last 50 iterations, but no full convergence was apparent (judged on the decreasing improvement in log-likelihood per iteration). Therefore, a second package, Wombat (Version 27-05-2020; Meyer, 2007), was tested with the same model and data. This package could only use the AI-REML algorithm, due to a residual covariance between repeated records on the same animal. It failed to find a maximum likelihood solution. Subsequently, a more parsimonious model was fitted by reducing the rank of fit for the genetic (co-)variance matrix from 10 to 9. This reduced model gave full convergence and a better log-likelihood compared with the log-likelihood of the full fit model (which did not fully converge). We compared the outcomes of the initial bivariate models, the incompletely converged ASReml and Wombat 10-trait models, and the converged reduced rank Wombat model. Heritabilities and genetic and phenotypic correlations differed only

slightly. Somewhat surprisingly, the SE of the estimate of the variance components for DMI increased unrealistically (compared with the univariate analysis). In the Results section we use the estimates of the bivariate analysis, and in Appendix 2, Table A1, we present the correlations estimated with the reduced-rank Wombat model.

RESULTS

Descriptive Statistics

Descriptive statistics for the recorded traits and for the 4 defined traits FIC, FS, RRI, and RCI are summarized in Table 2. The mean of FIC was 1.66 kg of DM/d, which means that the cows in the data set on average ate 1.66 kg of DM per day more than predicted by the simplification of the dairy cow model used by Zom (2014). The mean of FS, RRI, and RCI was 0 by construction, because these traits were defined as the residuals of a linear regression. The SD was largest for FIC and smallest for RCI. A similar trend was observed for DMI total, DMIR, and DMIC. The histograms in Figure A1 of Appendix 1 suggest that all traits follow a normal distribution. The strongest Pearson correlation (-0.84) was found between FS and RRI, but FS was moderately correlated with RCI (-0.56). The weakest Pearson correlation was found between FIC and RCI and between RRI and RCI, both below 0.11. Not only was the association between FIC and RCI, but it showed some underlying structure variation as well, probably due to different feeding systems in some experiments (Figure 1).

Table 2. Mean, minimum (min), maximum (max), and SD for all traits

Variable ¹	Mean	Min	Max	SD
Parity	2.76	1.00	11	1.54
DIM	99	2.00	365	73.2
MY (kg/d)	32.54	5.34	67.15	9.08
FPY (kg/d)	2.51	0.64	7.53	0.67
LW (kg)	637	306	932	77.97
DMI (kg/d)	21.15	8.00	35.00	4.13
DMIR (kg/d)	14.28	1.28	28.00	3.43
DMIC (kg/d)	6.94	0.00	25.57	2.86
Defined traits				
FIC (kg/d)	1.66	-11.44	14.39	3.66
FS (kg/d)	0.00	-7.97	7.95	2.29
RRI (kg/d)	0.00	-6.73	6.71	2.00
RCI (kg/d)	0.00	-4.91	4.91	1.40

¹MY = milk yield, FPY = sum of fat yield and protein yield, LW = live weight, DMIC = concentrate DMI, DMIR = roughage DMI, FIC = feed intake capacity, FS = feed saved, RRI = residual roughage intake, RCI = residual concentrate intake.

Relationship Between Feed Defined Traits and Other Traits

We found that FIC and RRI had a positive relationship with LW (Figure 2), and because of the change in sign FS had a negative association with LW. The estimated linear regression of the 4 defined traits on LW, MY, and DMI are presented in Table 3, adjusted for the fixed effects and therefore structural differences in the mean of experiments. The slope of FS on LW was after adjustment -0.0089 (Table 3), meaning that for every extra 100 kg of LW, the animal will consume an extra 0.89 kg of DM feed, which is within the range previously described by Ingvarsten (1994).

Of the 4 defined traits, only FIC had a positive association with milk production level (Figure 3 and Table 3), whereas FS, RRI, and RCI were not associated with milk production level, by construction.

The linear regression coefficients for all defined traits were associated with DMI (Table 3), and the regression of FIC on DMI was 1, illustrating that every kilogram of extra DMI also gave 1 kg additional FIC. Also, associations were linear and consistent across the range of DMI, except for the association between RCI and

DMI (Figure 4). At high DMI, the slope was steeper than at low or medium DMI, reflecting that cows with similar yields and roughage intakes must have a higher concentrate intake when DMI intake is high.

Multiparous cows showed, in general, a higher FIC, and first-parity cows had better efficiency (higher FS and lower RRI and RCI) than multiparous cows (Figure 5). For multiparous cows (parity 4+), RCI was more favorable than for cows in parity 2 or 3, which was not observed for FS and RRI.

For FIC a different pattern in relation to weeks in lactation was observed compared with the other defined traits (Figure 6). We found that FIC increased less steeply in early lactation and declined after 20 wk in milk. Values for RRI and RCI were low in the first weeks of lactation and rapidly increased, and a similar pattern was observed for FS. After 10 wk, values for RCI declined slowly, but RRI and FS stayed the same throughout lactation.

Genetic Analysis

Heritability and genetic correlations were estimated to analyze the genetic relationships among the 4 new

Figure 1. Phenotypic relationships between the 4 feed-defined traits. Each dot represents 1 record. On the left are the traits that represent the y-axis of each figure; at the bottom are the traits that represent the x-axis of each figure. FIC = feed intake capacity, FS = feed saved, RRI = residual roughage intake, RCI = residual concentrate intake, r = Pearson correlation coefficient.

Figure 2. Phenotypic relationships between live weight (kg) and FIC (A), FS (B), RRI (C), and RCI (D) in kilograms of DM/day. Each dot represents the mean of a certain trait for all records in a certain live weight class. Error bars indicate SE. FIC = feed intake capacity, FS = feed saved, RRI = residual roughage intake, RCI = residual concentrate intake.

defined traits and the underlying recorded traits. The largest h^2 (0.69) of all traits analyzed was found for LW (Table 4). We found that FIC had the highest heritability (0.19) and the largest phenotypic variance of the newly defined traits, whereas RCI and DMIC had the lowest h^2 (<0.04) of all traits. Heritability estimates of FIC, FS, and RRI were of similar sizes compared with heritability estimates of the milk production traits. Permanent environmental variance was highest for the MY traits and lowest for LW.

Table 5 presents the genetic and phenotypic correlations between all traits. The new trait FIC had genetic correlations close to unity with DMI and DMIR, and genetic correlations above 0.81 with the other new traits FS and RRI (in absolute terms). However, the genetic correlation of FIC and RCI was -0.19 . Moreover, a

strong genetic correlation of -0.96 was found between FS and RRI, suggesting that FS and RRI are genetically similar traits. Between FS and RCI, and RRI and RCI we detected little or no genetic correlation, which is in line with the results in Figure 1, where FS, RRI, and RCI showed dependence phenotypically, but RCI was distinct from RRI.

The genetic correlations between RRI and all other traits ranged from -0.07 (DMIC) to 0.91 (DMIR). We found that FIC, FS, and RRI showed very similar genetic correlations with LW: around 0.58 (in absolute terms). However, the genetic correlation between RCI and LW was estimated to be 0.03 ± 0.08 . These results were in line with the phenotypic relationships between these traits (Figure 2). Genetic correlations between RCI and milk production traits were negative (-0.20

Table 3. The linear regression of 4 feed-defined traits on live weight, milk yield, and DMI, with SE

Trait ¹	Live weight		Milk yield		DMI	
	Regression coefficient	SE	Regression coefficient	SE	Regression coefficient	SE
FIC	0.0171	0.0004	0.1570	0.0031	0.9994	0.0025
FS	-0.0089	0.0003	0.0200	0.0026	-0.7528	0.0029
RRI	0.0095	0.0003	-0.0078	0.0024	0.5729	0.0031
RCI	0.0004	0.0002	-0.0105	0.0016	0.2455	0.0027

¹FIC = feed intake capacity, FS = feed saved, RRI = residual roughage intake, RCI = residual concentrate intake.

Figure 3. Phenotypic relationships between milk production (kg/d) and FIC (A), FS (B), RRI (C), and RCI (D) in kilograms of DM/day. Each dot represents the mean value for a certain trait for all records in the same milk production class. Error bars indicate SE. FIC = feed intake capacity, FS = feed saved, RRI = residual roughage intake, RCI = residual concentrate intake.

to -0.30), whereas genetic correlations for FS and RRI were estimated to be unfavorable with the milk production traits. Even though FS, RRI, and RCI were phenotypically independent of milk production traits (Figure 3), they were genetically correlated.

DISCUSSION

Breeding for efficient cows is important for economic and environmental reasons. To breed for feed efficiency, a breeding goal trait needs to be defined, and scarce available data must be used optimally. It is expected that the amount of roughage in the ration of dairy cows may increase in the future, to reduce competition with human-edible food. Current breeding tools, such as RFI and FS, do not make a distinction between roughage and concentrate intake, although in some of our data a partial mixed ration is fed, with the concentrate portion recorded separately. Therefore, the aim of this research was to define alternative traits from our existing records that explain variation in FIC and roughage or concentrate intake separately and that quantify the phenotypic and genetic associations between these different trait definitions. Insights in these associations might help to judge whether it is worthwhile to breed for FIC or concentrate- and roughage-efficient cows separately.

Feed Intake Capacity

The mean of FIC was 1.66 kg of DM/d. This means that the cows in the data set on average ate 1.66 kg of DM extra per day. Feed intake capacity was based on a simplification of the dairy cow model (Zom, 2014). That our cows ate on average 1.66 kg more than predicted might be due to the data set used here versus that used by (Zom, 2014) to define and estimate the model parameters. The milk production level of the cows in our data set was higher, and the cows were on average 50 kg heavier. Also, some of the differences found here might also be because we were not able to model the pregnancy properly, due to lack of individual cow information on the day of conception.

Although nutritional models suggest that we should select cows that eat more roughage per kilogram of milk, and therefore also convert more grass into milk (Groen and Korver, 1989; Harrison et al., 1990; Veerkamp, 1998; Buckley et al., 2005), it is difficult to increase FIC as an independent trait. The genetic correlation of 0.98 between FIC and DMI, and of 0.89 between FIC and RRI, indicated that these traits are difficult to disentangle, which is not a surprise, considering that we measure DMI only and FIC is derived from this measurement only. No additional trait variation is included. At the same time, FIC is unfavorably correlated with

Figure 4. Phenotypic relationships between DMI (kg/d) and FIC (A), FS (B), RRI (C), and RCI (D) in kilograms of DM/day. Each dot represents the mean of a certain trait for all records in a DMI class. Error bars indicate SE. FIC = feed intake capacity, FS = feed saved, RRI = residual roughage intake, RCI = residual concentrate intake.

FS (-0.81) and LW (-0.58). Therefore, animals that have a high FIC eat more but process feed ingredients inefficiently or use them for purposes other than milk production.

Even when traits are correlated, using a selection index, genetic improvement for both traits may be possible (Brascamp, 1984). However, there is little room to improve both FIC and FS simultaneously, given the large unfavorable genetic correlation. Veerkamp and Koenen (1999) discussed the different options combining yield, LW, and DMI in a selection index, without trying to separate FIC from DMI, and demonstrated the expected influence of the different indices on yield, LW, and energy balance (which is similar in calculation to FS). They also concluded that changing the direction of selection of DMI (in combination with yield and LW) merely affects energy balance (or FS) and therefore depends on the genetic association with health and fertility. Our attempt in this study to separate FIC from DMI, and therefore allow for selection of cows that can process more roughage, did not change the conclusion that increasing FIC merely affects FS, as DMI and FIC are still strongly correlated.

In addition, Veerkamp et al. (2018) investigated the hypothesis that the extra weight and intake capacity from selection for a higher intake would be advanta-

geous on a lower-concentrate diet with more roughage. Selection index calculations demonstrated that selection for higher feed intake (without calculating FIC) resulted in heavier cows. When selecting for a higher feed intake relative to MY on a higher-density diet, no benefit were found in terms of profit on lower-density diets. Therefore, Veerkamp et al. (2018) concluded that breeding goals should always be for profit (fat- and protein-corrected milk production – DMI) independently of the nutritional environment, because of the absence of considerable $G \times E$. The definition of FIC and the estimated genetic parameters in this study are unlikely to change that conclusion and make selection for increased feed intake capacity useful, because FIC had a genetic correlation of 0.98 with DMI.

Feed Efficiency Traits

Feed efficiency by itself is a complex quantitative trait. Based on both theory (Kennedy et al., 1993) and data, the definition of feed efficiency can have a large influence on the selection response for the component traits. Hurley et al. (2017) showed that genetic correlations among different definitions of feed efficiency ranged between -0.8 and 1 . We used FS as a measure of efficiency (Pryce et al., 2015), and FS showed genetic

Figure 5. Relationships between parity and FIC (A), FS (B), RRI (C), and RCI (D). All traits given in kilograms of DM/day extra. Mean per parity is shown for all traits. Vertical bars indicate SE. FIC = feed intake capacity, FS = feed saved, RRI = residual roughage intake, RCI = residual concentrate intake.

variation and heritability of 0.10 ± 0.01 (mean $\pm$ SE). The difference between FS and RFI is that energy sinks due to LW were not taken into account, and therefore FS and LW were correlated phenotypically as well as genetically. The genetic correlation of -0.58 between LW and FS demonstrated that approximately a quarter of the genetic variance in FS can be explained by LW, and the negative direction of the correlation shows that genetically heavier cows had lower values for FS (i.e., were less efficient) because of feed used for additional maintenance. In the literature, estimation of heritability for RFI traits ranged from 0.01 to 0.36 (Coleman et al., 2010; Vallimont et al., 2011; Connor et al., 2013; Tempelman et al., 2015; Manafiazar et al., 2016). However, the definition of RFI is often very different in literature and especially the measurements used, daily or averaged across a period, and the models used, from which the residual is taken, differ (Berry and Crowley, 2013; Berry and Pryce, 2014). Therefore, any measure of efficiency, such as FS or RFI, must be discussed within the context of the definition and data used and not as an existing trait.

Phenotypic and genetic correlations between FS and DMI reflect the proportion of variance in DMI not explained by the component traits in the model. Correla-

tions presented here were similar to those previously reported between RFI and DMI (Veerkamp et al., 1995; Manzanilla Pech et al., 2014). Although the phenotypic and genetic relationship between DMI and FS is obvious, different biological processes can influence covariation in DMI and FS. Larger meal sizes have a negative effect on digestibility (de Souza et al., 2018), and Potts et al. (2017) suggested that differences in digestibility may influence divergence for RFI in cows. However, they hint that other factors, such as metabolic processes and heat production, may have a larger effect on FS. It has also been reported that high-efficiency cows had lower DMI and lower peak meal sizes, spent more time ruminating per kilogram of DM, and showed higher digestibility than low-efficiency cows (Ben Meir et al., 2018). However, it is still not clear to what extent biological processes cause differences in FS. As defined here, FS showed a relationship to DMI similar to that of FS as defined in literature.

Our results showed that in the first 5 wk of lactation, feed efficiency measures were negative (Figure 6). Low-RRI or -RCI and high-FS animals could be mobilizing tissue reserves to maintain milk production. Hence, they eat less than expected based on milk production (Pryce et al., 2014). It has been found that cows with

Figure 6. Relationship between DIM (in bins per week) and FIC (A), FS (B), RRI (C), and RCI (D). Mean per week in milk is shown for all traits. All traits given in kilograms of DM/day extra. Vertical bars indicate SE. FIC = feed intake capacity, FS = feed saved, RRI = residual roughage intake, RCI = residual concentrate intake.

low RFI are more likely to be in negative energy balance than less-efficient cows (Seymour et al., 2020). Negative energy balance is associated with impaired health and fertility (Collard et al., 2000; Walsh et al., 2011). However, no clear relationship with feed efficiency and fertility has been found (Coleman et al., 2010; Olson et al., 2011). Genetic correlations with health and fertility traits must be considered in breeding goals (Veerkamp et al., 2013; VandeHaar et al., 2016). Estimates of ge-

netic correlations between efficiency traits and fertility are available, but they have large SE (Vallimont et al., 2013). To avoid unfavorable correlated responses, reliable estimates of genetic correlations between efficiency, health, and fertility must be obtained.

Clear mean differences for efficiency measures exist across parity and lactation stage (Figures 5 and 6). For breeding, it is important that traits are repeatable across life stages and environments (VandeHaar et al.,

Table 4. The phenotypic variance (σ_p^2), permanent environmental variance (σ_{perm}^2), and genetic variance (σ_a^2), and the repeatability ($p^2 = \sigma_{perm}^2 / \sigma_p^2$) and heritability ($h^2 = \sigma_a^2 / \sigma_p^2$) for all traits, with SE

Trait ¹	σ_p^2	σ_{perm}^2	σ_a^2	p^2	SE	h^2	SE
MY	29.0	14.5	6.7	0.50	0.016	0.23	0.018
FPY	0.18	0.08	0.02	0.46	0.014	0.14	0.016
LW	3,041	554	2,104	0.18	0.008	0.69	0.010
DMI	6.03	2.04	0.95	0.34	0.012	0.16	0.014
DMIC	2.36	0.91	0.09	0.39	0.011	0.04	0.009
DMIR	4.52	1.22	0.85	0.27	0.011	0.19	0.013
FIC	7.19	2.06	1.35	0.28	0.011	0.19	0.014
FS	4.26	1.00	0.44	0.24	0.010	0.10	0.011
RRI	3.44	0.77	0.48	0.22	0.010	0.14	0.012
RCI	1.60	0.42	0.04	0.26	0.008	0.03	0.007

¹MY = milk yield, FPY = sum of fat yield and protein yield, LW = live weight, DMIC = concentrate DMI, DMIR = roughage DMI, FIC = feed intake capacity, FS = feed saved, RRI = residual roughage intake, RCI = residual concentrate intake.

Table 5. Phenotypic¹ (above diagonal) and genetic correlations (below diagonal) between the traits in the bivariate analysis, with SE in parentheses

Trait ²	MY	FPY	LW	DMI	DMIC	DMIR	FIC	FS	RRI	RCI
MY	1									
FPY	0.61 (0.04)	0.74								
LW	0.13 (0.04)	0.23 (0.05)	1							
DMI	0.61 (0.04)	0.76 (0.04)	0.58 (0.04)	1						
DMIC	0.52 (0.08)	0.38 (0.11)	0.17 (0.08)	0.36 (0.09)	1					
DMIR	0.47 (0.05)	0.67 (0.05)	0.57 (0.03)	0.95 (0.01)	0.07 (0.11)	1				
FIC	0.55 (0.04)	0.71 (0.04)	0.58 (0.03)	0.98 (0.01)	0.13 (0.10)	0.99 (0.00)	1			
FS	-0.14 (0.07)	-0.33 (0.08)	-0.58 (0.03)	-0.79 (0.03)	-0.04 (0.12)	-0.99 (0.00)	-0.81 (0.03)	1		
RRI	0.18 (0.06)	0.39 (0.07)	0.58 (0.03)	0.84 (0.03)	-0.07 (0.11)	0.91 (0.02)	0.89 (0.02)	-0.96 (0.01)	1	
RCI	-0.20 (0.11)	-0.30 (0.13)	0.03 (0.08)	-0.04 (0.11)	0.58 (0.11)	-0.22 (0.10)	-0.19 (0.10)	0.25 (0.10)	0.00 (0.11)	1

¹SE for phenotypic correlations ≤ 0.01 .²MY = milk yield, FPY = sum of fat yield and protein yield, LW = live weight, DMIC = concentrate DMI, DMIR = roughage DMI, FIC = feed intake capacity, FS = feed saved, RRI = residual roughage intake, RCI = residual concentrate intake.

2016). It has previously been shown that feed efficiency measures fluctuate throughout lactation (de Haas et al., 2017; Hurley et al., 2017) and that feed efficiency in one parity is not the same trait as feed efficiency in another parity (Hurley et al., 2018). By contrast, VandeHaar et al. (2016) argue that RFI is repeatable over lactation stages. Li et al. (2017) recommend taking lactation stage into account when assessing RFI. Connor et al. (2019) investigated when and how often RFI should be measured to accurately predict RFI for the whole lactation. Lactation stage and parity should be considered when assessing feed efficiency, whatever trait definition is used.

Concentrate Versus Roughage Efficiency

We defined RRI and RCI to see whether differences in variation exist between concentrate and roughage efficiency. Based on the phenotypic and genetic correlations, RRI and FS were found to be similar traits. The genetic correlation between RRI and DMIR was large (0.91), suggesting that 81% of the variation in DMIR was not explained by concentrate intake and the milk component traits. In comparison, the genetic correlation between FS and DMI was -0.79 ; thus, 64% of the variation was not explained by yield alone. Genetic correlation with concentrate intake was close to zero for FS. Residual concentrate intake showed the smallest genetic variation and the lowest heritability of our defined traits, and the relationships with the other traits were different for RCI than for FS and RRI. Residual concentrate intake was independent of LW and had negative genetic correlations with milk production traits, in the direction that when RCI decreased (i.e., more efficient cows), milk production increased. Hence, although RCI was a phenotypically distinct trait from FS and RRI, heritability for DMIC and RCI was virtually zero in our data set. Furthermore the genetic correlation between DMI and DMIR was 0.96, suggesting that roughage intake explained nearly all of the genetic variation in total feed intake in our data. These results agree with the results of a Swedish study, which found that the genetic correlation between forage intake and total intake was above 0.90 for Holstein and 0.83 for Swedish Red during the lactation period (Tarekegn et al., 2021). Obviously, this genetic correlation always depends on concentrate feeding system. In our data there was no fixed concentrate gift protocol across the experiments, and in about 10% of the experimental treatments all cows received the same amount of concentrate and ate all the concentrate given. In other experiments, concentrate gift might have depended on lactation stage or yield level. Therefore the most likely reason that (residual) concentrate intake had such a

low heritability is that concentrate intake is primarily determined by the feeding system and not by the genetics of the cow.

Despite the definition and the name of RRI, we must ask whether selecting for RRI really selects animals that are better suited to convert roughage into milk. This is especially important because concentrate and roughage intake are not independent (Faverdin et al., 1991). In our definition of RRI, it could still be that animals produce more efficiently from concentrate and therefore require less roughage at the same yield and concentrate levels. For cows with a high RRI, 27% of DMI consisted of concentrate. For cows with a low RRI (i.e., efficient with roughage), almost 40% of the DMI consisted of concentrate. Hence, low-RRI cows ate less roughage compared with high-RRI cows. For cows with a low RCI (i.e., efficient with concentrate), 24% of the DMI consisted of concentrate. Thus, low-RCI cows processed more roughage than low-RRI cows. Although these scale effects are clear across the diets, it is unclear how they play a role when comparing animals after adjusting for experimental treatments. More sophisticated adjustments might be required to fully separate concentrate and roughage intake, and it is important to investigate the biological differences between animals for these traits.

In this research, roughage and concentrate intake were separated, and for roughage intake the total intake of the partial mixed ration was used. From a nutritional perspective this is suboptimal. Intake of concentrate limits the intake of roughage (Faverdin et al., 1991). Additionally, there is huge variety in types of roughage and concentrate. For example, roughage may be either straw or maize silage. These types of roughages have different amounts of fiber, starch, and protein. Moreover, many different types of concentrate exist, but this information was not recorded for all data. The distinction between concentrate and roughage made in this study was a beginning and gave the opportunity to investigate differences between roughage and concentrate efficiency. Classification on a continuous scale (e.g., based on yield response per kg of DMI) might be another way to classify diets (Veerkamp et al., 2018).

CONCLUSIONS

This research aimed to define alternative traits that differentiate between FIC and roughage or concentrate intake, and to investigate the phenotypic and genetic relationships among them. We found that FIC was strongly correlated with DMI. The RCI trait was distinct from other definitions; however, RCI was hardly heritable. The traits FS and RRI were found to be genetically similar. Therefore, selection for FIC

is expected to simply increase DMI and RRI. There seems to be little advantage in separating concentrate and roughage intake, or to select for FIC in our data, because measured concentrate intake is determined by the feeding system and not by the genetics of the cow.

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REFERENCES

- Arthur, P. F., and R. M. Herd. 2005. Efficiency of feed utilisation by livestock—Implications and benefits of genetic improvement. *Can. J. Anim. Sci.* 85:281–290. <https://doi.org/10.4141/A04-062>.
- Beerda, B., W. Ouweltjes, L. B. J. Šebek, J. J. Windig, and R. F. Veerkamp. 2007. Effects of genotype by environment interactions on milk yield, energy balance, and protein balance. *J. Dairy Sci.* 90:219–228. [https://doi.org/10.3168/jds.S0022-0302\(07\)72623-1](https://doi.org/10.3168/jds.S0022-0302(07)72623-1).
- Ben Meir, Y. A., M. Nikbachat, Y. Fortnik, S. Jacoby, H. Levit, G. Adin, M. Cohen Zinder, A. Shabtay, E. Gershon, M. Zachut, S. J. Mabjeesh, I. Halachmi, and J. Miron. 2018. Eating behavior, milk production, rumination, and digestibility characteristics of high- and low-efficiency lactating cows fed a low-roughage diet. *J. Dairy Sci.* 101:10973–10984. <https://doi.org/10.3168/jds.2018-14684>.
- Berry, D. P., and J. J. Crowley. 2013. Cell Biology Symposium: Genetics of feed efficiency in dairy and beef cattle. *J. Anim. Sci.* 91:1594–1613. <https://doi.org/10.2527/jas.2012-5862>.
- Berry, D. P., and J. E. Pryce. 2014. Feed efficiency in growing and mature animals. *Proc. 10th World Congress of Genetics Applied to Livestock Production, Vancouver, Canada*. Page 109 in *Genetics of Trait Complexes: Feed Intake and Efficiency*.
- Brascamp, E. W. 1984. Selection indices with constraints. *Animal Breeding Abstracts* 52:645–654.
- Buckley, F., C. Holmes, and M. G. Keane. 2005. Genetic characteristics required in dairy and beef cattle for temperate grazing systems. Pages 61–78 in *Utilisation of Grazed Grass in Temperate Animal Systems. Proc. Satellite Workshop of the XXth International Grassland Congress, Cork, Ireland*. J. J. Murphy, ed.
- Coffey, M. P., G. Simm, J. D. Oldham, W. G. Hill, and S. Brotherstone. 2004. Genotype and diet effects on energy balance in the first three lactations of dairy cows. *J. Dairy Sci.* 87:4318–4326. [https://doi.org/10.3168/jds.S0022-0302\(04\)73577-8](https://doi.org/10.3168/jds.S0022-0302(04)73577-8).
- Coleman, J., D. P. Berry, K. M. Pierce, A. Brennan, and B. Horan. 2010. Dry matter intake and feed efficiency profiles of 3 genotypes of Holstein-Friesian within pasture-based systems of milk production. *J. Dairy Sci.* 93:4318–4331. <https://doi.org/10.3168/jds.2009-2686>.
- Collard, B. L., P. J. Boettcher, J. C. M. Dekkers, D. Petitclerc, and L. R. Schaeffer. 2000. Relationships between energy balance and health traits of dairy cattle in early lactation. *J. Dairy Sci.* 83:2683–2690. [https://doi.org/10.3168/jds.S0022-0302\(00\)75162-9](https://doi.org/10.3168/jds.S0022-0302(00)75162-9).
- Connor, E. E., J. L. Hutchison, H. D. Norman, K. M. Olson, C. P. Van Tassell, J. M. Leith, and R. L. Baldwin. 2013. Use of residual feed intake in Holsteins during early lactation shows potential to improve feed efficiency through genetic selection. *J. Anim. Sci.* 91:3978–3988. <https://doi.org/10.2527/jas.2012-5977>.
- Connor, E. E., J. L. Hutchison, C. P. Van Tassell, and J. B. Cole. 2019. Defining the optimal period length and stage of growth or lacta-

- tion to estimate residual feed intake in dairy cows. *J. Dairy Sci.* 102:6131–6143. <https://doi.org/10.3168/jds.2018-15407>.
- CRV. 2019. Jaarstatistieken 2018 voor Nederland [Annual Statistics 2018 for the Netherlands]. Cooperatie CRV.
- CVB. 2019. Veevoedertabel 2019 [Animal feed table 2019]. In *Chemische samenstellingen en nutritionele waarden van voedermiddelen* [Chemical Compositions and Nutritional Values of Feed Materials]. Federatie Nederlandse Diervoederketen.
- de Haas, Y., M. Pszczola, H. Soyeurt, E. Wall, and J. Lassen. 2017. Invited review: Phenotypes to genetically reduce greenhouse gas emissions in dairying. *J. Dairy Sci.* 100:855–870. <https://doi.org/10.3168/jds.2016-11246>.
- de Souza, R. A., R. J. Tempelman, M. S. Allen, W. P. Weiss, J. K. Bernard, and M. J. VandeHaar. 2018. Predicting nutrient digestibility in high-producing dairy cows. *J. Dairy Sci.* 101:1123–1135. <https://doi.org/10.3168/jds.2017-13344>.
- FADN. 2011. Farm Economics Brief. EU Production Costs Overview. European Commission–EU FADN. Accessed Feb. 20, 2021. <https://ec.europa.eu/agriculture/rica/pdf/Brief201102.pdf>.
- Faverdin, P., J. P. Dulphy, J. B. Coulon, R. Vêrité, J. P. Garel, J. Rouel, and B. Marquis. 1991. Substitution of roughage by concentrates for dairy cows. *Livest. Prod. Sci.* 27:137–156. [https://doi.org/10.1016/0301-6226\(91\)90092-5](https://doi.org/10.1016/0301-6226(91)90092-5).
- Gilmour, A. R., B. Gogel, B. Cullis, S. Welham, and R. Thompson. 2015. ASReml User Guide Release 4.1 Functional Specification. VSN International Ltd.
- Groen, A. F., and S. Korver. 1989. The economic value of feed intake capacity of dairy cows. *Livest. Prod. Sci.* 22:269–281. [https://doi.org/10.1016/0301-6226\(89\)90060-2](https://doi.org/10.1016/0301-6226(89)90060-2).
- Gunsett, F. C. 1984. Linear index selection to improve traits defined as ratios. *J. Anim. Sci.* 59:1185–1193. <https://doi.org/10.2527/jas1984.5951185x>.
- Harrison, R. O., S. P. Ford, J. W. Young, A. J. Conley, and A. E. Freeman. 1990. Increased milk production versus reproductive and energy status of high producing dairy cows. *J. Dairy Sci.* 73:2749–2758. [https://doi.org/10.3168/jds.S0022-0302\(90\)78960-6](https://doi.org/10.3168/jds.S0022-0302(90)78960-6).
- Hilhorst, G. 2018. Eiwit van eigen land: Het begint met het goed benutten van gras [Protein from your own land: It starts with making good use of grass]. Vol. 2020. Accessed Feb. 20, 2021. www.kringloopwijzer.nl.
- Hurley, A. M., N. López-Villalobos, S. McParland, E. Lewis, E. Kennedy, M. O'Donovan, J. L. Burke, and D. P. Berry. 2017. Genetics of alternative definitions of feed efficiency in grazing lactating dairy cows. *J. Dairy Sci.* 100:5501–5514. <https://doi.org/10.3168/jds.2016-12314>.
- Hurley, A. M., N. López-Villalobos, S. McParland, E. Lewis, E. Kennedy, M. O'Donovan, J. L. Burke, and D. P. Berry. 2018. Characteristics of feed efficiency within and across lactation in dairy cows and the effect of genetic selection. *J. Dairy Sci.* 101:1267–1280. <https://doi.org/10.3168/jds.2017-12841>.
- Ingvarsen, K. L. 1994. Models of voluntary food intake in cattle. *Livest. Prod. Sci.* 39:19–38. [https://doi.org/10.1016/0301-6226\(94\)90149-X](https://doi.org/10.1016/0301-6226(94)90149-X).
- Kennedy, B. W., J. H. van der Werf, and T. H. Meuwissen. 1993. Genetic and statistical properties of residual feed intake. *J. Anim. Sci.* 71:3239–3250. <https://doi.org/10.2527/1993.71123239x>.
- Kleijne, J. d., S. Ruiter, A. v. Velde, H. Zonderland, K. Bolscher, N. Oerlemans, W. d. Weijden, L. Hendrix, J. J. v. Dijk, and T. Loman. 2018. Grondgebondenheid als basis voor een toekomstbestendige melkveehouderij [Land-relatedness as the basis for future-proof dairy farming]. De Commissie Grondgebondenheid [The Land-Relatedness Committee].
- Koch, R. M., L. A. Swiger, D. Chambers, and K. E. Gregory. 1963. Efficiency of feed use in beef cattle. *J. Anim. Sci.* 22:486–494. <https://doi.org/10.2527/jas1963.222486x>.
- Koenen, E. P. C., P. B. M. Berentsen, and A. F. Groen. 2000. Economic values of live weight and feed intake capacity of dairy cattle under Dutch production circumstances. *Livest. Prod. Sci.* 66:235–250. [https://doi.org/10.1016/S0301-6226\(00\)00167-6](https://doi.org/10.1016/S0301-6226(00)00167-6).
- Li, B., B. Berglund, W. F. Fikse, J. Lassen, M. H. Lidauer, P. Mäntysaari, and P. Løvendahl. 2017. Neglect of lactation stage leads to naive assessment of residual feed intake in dairy cattle. *J. Dairy Sci.* 100:9076–9084. <https://doi.org/10.3168/jds.2017-12775>.
- Li, B., P. M. VanRaden, E. Guduk, J. R. O'Connell, D. J. Null, E. E. Connor, M. J. VandeHaar, R. J. Tempelman, K. A. Weigel, and J. B. Cole. 2020. Genomic prediction of residual feed intake in US Holstein dairy cattle. *J. Dairy Sci.* 103:2477–2486. <https://doi.org/10.3168/jds.2019-17332>.
- Manafiazar, G., L. Goonewardene, F. Miglior, D. H. Crews Jr., J. A. Basarab, E. Okine, and Z. Wang. 2016. Genetic and phenotypic correlations among feed efficiency, production and selected conformation traits in dairy cows. *Animal* 10:381–389. <https://doi.org/10.1017/S1751731115002281>.
- Manzanilla Pech, C. I. V., R. F. Veerkamp, M. P. L. Calus, R. Zom, A. van Knegsel, J. E. Pryce, and Y. De Haas. 2014. Genetic parameters across lactation for feed intake, fat-and protein-corrected milk, and liveweight in first-parity Holstein cattle. *J. Dairy Sci.* 97:5851–5862. <https://doi.org/10.3168/jds.2014-8165>.
- Meyer, K. 2007. WOMBAT—A tool for mixed model analyses in quantitative genetics by restricted maximum likelihood (REML). *J. Zhejiang Univ. Sci. B* 8:815–821. <https://doi.org/10.1631/jzus.2007.B0815>.
- Olson, K. M., B. G. Cassell, M. D. Hanigan, and R. E. Pearson. 2011. Short communication: Interaction of energy balance, feed efficiency, early lactation health events, and fertility in first-lactation Holstein, Jersey, and reciprocal F₁ crossbred cows. *J. Dairy Sci.* 94:507–511. <https://doi.org/10.3168/jds.2010-3433>.
- Potts, S. B., J. P. Boerman, A. L. Lock, M. S. Allen, and M. J. VandeHaar. 2017. Relationship between residual feed intake and digestibility for lactating Holstein cows fed high and low starch diets. *J. Dairy Sci.* 100:265–278. <https://doi.org/10.3168/jds.2016-11079>.
- Pryce, J. E., O. Gonzalez-Recio, G. Nieuwhof, W. J. Wales, M. P. Coffey, B. J. Hayes, and M. E. Goddard. 2015. Hot topic: Definition and implementation of a breeding value for feed efficiency in dairy cows. *J. Dairy Sci.* 98:7340–7350. <https://doi.org/10.3168/jds.2015-9621>.
- Pryce, J. E., B. L. Nielsen, R. F. Veerkamp, and G. Simm. 1999. Genotype and feeding system effects and interactions for health and fertility traits in dairy cattle. *Livest. Prod. Sci.* 57:193–201. [https://doi.org/10.1016/S0301-6226\(98\)00180-8](https://doi.org/10.1016/S0301-6226(98)00180-8).
- Pryce, J. E., W. J. Wales, Y. de Haas, R. F. Veerkamp, and B. J. Hayes. 2014. Genomic selection for feed efficiency in dairy cattle. *Animal* 8:1–10. <https://doi.org/10.1017/S1751731113001687>.
- R Core Team. 2019. R: A Language and Environment for Statistical Computing. R Foundation for Statistical Computing.
- Seymour, D. J., A. Cánovas, T. C. S. Chud, J. P. Cant, V. R. Osborne, C. F. Baes, F. S. Schenkel, and F. Miglior. 2020. The dynamic behavior of feed efficiency in primiparous dairy cattle. *J. Dairy Sci.* 103:1528–1540. <https://doi.org/10.3168/jds.2019-17414>.
- Tarekegn, M., J. Karlsson, C. Kronqvist, B. Berglund, K. Holtenius, and E. Strandberg. 2021. Genetic parameters of forage dry matter intake and milk produced from forage in Swedish Red and Holstein dairy cows. *J. Dairy Sci.* 104:4424–4440. <https://doi.org/10.3168/jds.2020-19224>.
- Tempelman, R. J., and Y. Lu. 2020. Symposium review: Genetic relationships between different measures of feed efficiency and the implications for dairy cattle selection indexes. *J. Dairy Sci.* 103:5327–5345. <https://doi.org/10.3168/jds.2019-17781>.
- Tempelman, R. J., D. M. Spurlock, M. Coffey, R. F. Veerkamp, L. E. Armentano, K. A. Weigel, Y. de Haas, C. R. Staples, E. E. Connor, Y. Lu, and M. J. VandeHaar. 2015. Heterogeneity in genetic and nongenetic variation and energy sink relationships for residual feed intake across research stations and countries. *J. Dairy Sci.* 98:2013–2026. <https://doi.org/10.3168/jds.2014.8510>.
- Vallimont, J. E., C. D. Dechow, J. M. Daubert, M. W. Dekleva, J. W. Blum, C. M. Barlieb, W. Liu, G. A. Varga, A. J. Heinrichs, and C. R. Baumrucker. 2011. Short communication: Heritability of gross feed efficiency and associations with yield, intake, residual intake, body weight, and body condition score in 11 commercial Pennsylvania tie stalls. *J. Dairy Sci.* 94:2108–2113. <https://doi.org/10.3168/jds.2010-3888>.

- Vallimont, J. E., C. D. Dechow, J. M. Daubert, M. W. Dekleva, J. W. Blum, W. Liu, G. A. Varga, A. J. Heinrichs, and C. R. Baumrucker. 2013. Short communication: Feed utilization and its associations with fertility and productive life in 11 commercial Pennsylvania tie-stall herds. *J. Dairy Sci.* 96:1251–1254. <https://doi.org/10.3168/jds.2012-5712>.
- van Kneegsel, A. T. M., G. J. Rummelink, S. Jorjone, V. Fievez, and B. Kemp. 2014. Effect of dry period length and dietary energy source on energy balance, milk yield, and cow composition of dairy cows. *J. Dairy Sci.* 97:1499–1512. <https://doi.org/10.3168/jds.2013-7391>.
- VandeHaar, M. J., L. E. Armentano, K. Weigel, D. M. Spurlock, R. J. Tempelman, and R. Veerkamp. 2016. Harnessing the genetics of the modern dairy cow to continue improvements in feed efficiency. *J. Dairy Sci.* 99:4941–4954. <https://doi.org/10.3168/jds.2015-10352>.
- Veerkamp, R. F. 1998. Selection for economic efficiency of dairy cattle using information on live weight and feed intake: A review. *J. Dairy Sci.* 81:1109–1119. [https://doi.org/10.3168/jds.S0022-0302\(98\)75673-5](https://doi.org/10.3168/jds.S0022-0302(98)75673-5).
- Veerkamp, R. F., and G. C. Emmans. 1995. Sources of genetic variation in energetic efficiency of dairy cows. *Livest. Prod. Sci.* 44:87–97. [https://doi.org/10.1016/0301-6226\(95\)00065-0](https://doi.org/10.1016/0301-6226(95)00065-0).
- Veerkamp, R. F., G. C. Emmans, A. R. Cromie, and G. Simm. 1995. Variance components for residual feed intake in dairy cows. *Livest. Prod. Sci.* 41:111–120. [https://doi.org/10.1016/0301-6226\(94\)00056-D](https://doi.org/10.1016/0301-6226(94)00056-D).
- Veerkamp, R. F., and E. P. C. Koenen. 1999. Genetics of food intake, live weight, condition score and energy balance. BSAP Occasional Publication 24:63–73. <https://doi.org/10.1017/S1463981500043089>.
- Veerkamp, R. F., J. K. Oldenbroek, H. J. Van der Gaast, and J. H. Van der Werf. 2000. Genetic correlation between days until start of luteal activity and milk yield, energy balance, and live weights. *J. Dairy Sci.* 83:577–583. [https://doi.org/10.3168/jds.S0022-0302\(00\)74917-4](https://doi.org/10.3168/jds.S0022-0302(00)74917-4).
- Veerkamp, R. F., J. E. Pryce, D. Spurlock, D. P. Berry, M. Coffey, P. Løvendahl, R. van der Linde, J. Bryant, F. Miglior, Z. Wang, M. Winters, N. Krattenmacher, N. Charfeddine, J. Pedersen, and Y. de Haas. 2013. Selection on feed intake or feed efficiency: A position paper from gDMI breeding goal discussions. Pages 15–22 in *Proc. Interbull Meeting No. 47*, Nantes, France. Interbull.
- Veerkamp, R. F., G. C. B. Schopen, and Y. de Haas. 2018. Selection for low or high feed intake cows: Genotype by environment interaction for milk yield, live weight and dry matter intake in dairy cows. Page 189 in *Proc. World Congress on Genetics Applied to Livestock Production*. Auckland, New Zealand.
- Walsh, S. W., E. J. Williams, and A. C. O. Evans. 2011. A review of the causes of poor fertility in high milk producing dairy cows. *Anim. Reprod. Sci.* 123:127–138. <https://doi.org/10.1016/j.anireprosci.2010.12.001>.
- Zom, R. 2014. The development of a model for the prediction of feed intake and energy partitioning in dairy cows. PhD thesis. Animal Nutrition, Wageningen Institute of Animal Sciences, Wageningen, the Netherlands.
- Zom, R. L. G., G. André, and A. M. van Vuuren. 2012. Development of a model for the prediction of feed intake by dairy cows: 1. Prediction of feed intake. *Livest. Sci.* 143:43–57. <https://doi.org/10.1016/j.livsci.2011.08.014>.

ORCIDS

- Ghyslaine C. B. Schopen <https://orcid.org/0000-0003-3941-9872>
 Birgit Gredler-Grandl <https://orcid.org/0000-0002-5916-8639>
 Roel F. Veerkamp <https://orcid.org/0000-0002-5240-6534>

APPENDIX 1

Figure A1. Histograms of all traits in this study. DMIR = roughage DMI, DMIC = concentrate DMI, FIC = feed intake capacity, FS = feed saved, RRI = residual roughage intake, RCI = residual concentrate intake.

APPENDIX 2

Table A1. Phenotypic¹ (above diagonal) and genetic correlations (below diagonal) between the traits in a reduced-rank multitrait analysis, with SE in parentheses

Trait ²	MY	FPY	LW	DMI	DMIC	DMIR	FIC	FS	RRI	RCI
MY	1									
FPY	0.63 (0.04)	1								
LW	0.14 (0.04)	0.24 (0.05)	1							
DMI	0.59 (0.11)	0.72 (0.04)	0.56 (0.12)	1						
DMIC	0.48 (0.09)	0.38 (0.10)	0.12 (0.08)	0.41 (0.31)	1					
DMIR	0.48 (0.05)	0.66 (0.04)	0.57 (0.03)	0.96 (0.13)	0.12 (0.10)	1				
FIC	0.54 (0.04)	0.68 (0.04)	0.58 (0.03)	0.98 (0.00)	0.23 (0.10)	0.99 (0.00)	1			
FS	-0.17 (0.06)	-0.31 (0.07)	-0.57 (0.03)	-0.82 (0.11)	-0.14 (0.11)	-0.85 (0.03)	-0.84 (0.03)	1		
RRI	0.23 (0.06)	0.40 (0.06)	0.59 (0.03)	0.84 (0.16)	0.01 (0.10)	0.91 (0.02)	0.89 (0.02)	-0.97 (0.01)	1	
RCI	-0.22 (0.11)	-0.35 (0.12)	0.00 (0.08)	0.00 (0.40)	0.52 (0.12)	-0.16 (0.10)	-0.10 (0.10)	0.25 (0.10)	0.03 (0.11)	1

¹SE for phenotypic correlations <0.04.

²MY = milk yield, FPY = sum of fat yield and protein yield, LW = live weight, DMIC = concentrate DMI, DMIR = roughage DMI, FIC = feed intake capacity, FS = feed saved, RRI = residual roughage intake, RCI = residual concentrate intake.